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Appendix 1

PARAMETERS OF THE TWO AREA POWER SYSTEM INVESTIGATED

(a)	Parameters of the thermal system						
	Nominal frequency		= 50 Hz				
	Rated power of each area P_{r_1}, P_{r_2}		= 2000 1	MW			
•	Nominal load of each area		= 1000 !	MW			
	Tie line synchronising coefficient T_{12}		= 0.25 j	pu			
	Equivalent load frequency droop R		= 2.0 H	z/puMW			
·	Steam governor-turbine data:						
	Governor actuator time constant T_{G}		= 0.1 s				
	Turbine time constant T _T		= 0.5 s				
(b)	University of Moratuwa, Sri Parameters of the mixed hydro-thermal s Electronic Theses & Disserta www.lib.mrt.ac.lk Percentage of regulating hydro units	Lanka. ^{ystem} luons	= 20%				
ŀ	Percentage of regulating thermal units		= 80%				
	Hydro governor-turbine data:						
	Governor actuator time constants:	T_1	= 40 s				
		T ₂	= 0.513	S			
		т _в	= 5.0 s				
	Penstock time constant	T,	= 1.0 s				

Rest of the data are as in (a) above

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Appendix 2

IDENTIFICATION OF THE PARAMETERS OF THE MODEL USING LEAST SQUARES ESTIMATION

The least squares estimation for identification of model parameters will be discussed and a recursive least squares estimation for on line parameter estimation will be derived.

A2.1 Least Squares Estimation

Consider the system model given by equation (A2.1)

$$y(t) = q^{-k} \frac{B}{A} u(t)$$
 (A2.1)

where,

and

 $\bigcup_{B} = \bigcup_{i=1}^{n} \bigcup_{j=1}^{n} \bigcup_{i=1}^{n} \bigcup_{i=1}^{n} \bigcup_{i=1}^{n} \bigcup_{j=1}^$

The coefficients of A and B are the unknown parameters to be estimated. Equation (A2.1) can be expressed in difference equation form as:

$$y(t) = -a_1y(t-1) - a_2y(t-2) \dots - a_n y(t-n) + b_0u(t-k) + b_1u(t-k-1) + \dots + b_m u(t-k-m) (A2.2)$$

where,

$$y(t)$$
, $y(t-1)$, $y(t-n)$, $u(t-k)$, $u(t-k-1)$ $u(t-k-m)$
are known output and input data.

and
$$a_1, a_2, \ldots, a_n, b_0, b_1, \ldots, b_m$$
 are unknown model parameters
to be estimated.

Let the number of unknown parameters to be N

Equation (A2.2) can be written for N sampling instances to obtain the formorized relations. $y(t-N+1) = -a_1 y(t-N) - \dots - a_n y(t-N+1-n) + b_0 u(t-N+1-k) + \dots + b_m u(t-N+1-k-m) + b_m u(t-N+1-k-m)$ $(t-1) = -a_1 y(t-2) - \dots - a_n y(t-1-n) + b_0 u(t-1-k) + \dots + b_m u(t-1-k-m)$ $y(t) = -a_1 y(t-1) - \dots - a_n y(t-n) + b_0 u(t-k) + \dots + b_m u(t-k-m)$

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The set of equations (A2.13) trepresent a set of N simultaneous linear equations with N unknowns, which can be solved to find the N unknown model parameters provided the N equations are independent.

Equation (A2.3) can be written in matrix form as:

$$\underline{Y} = X \underline{\theta}$$
(A2.4)

where Y is a vector consisting of the output data

i.e.
$$Y = (y(t-N+1), y(t-N+2), \dots, y(t-1), y(t))^T$$
 (A2.5)

 θ is the unknown parameter vector given by

$$\underline{\theta} = (-a_1, -a_2, \dots -a_n, b_0, b_1, \dots b_m)^T$$
$$= (\theta_1, \theta_2, \dots \theta_N)^T$$

and X is a (N,N) matrix consisting of past input-output data values:

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(A2.3)

	y(t-N)y(t-N+1-n)		u(t-n+1-k)	u(t-N+1-k-m)	
	•	•	•	•	
•	•	•	•	•	
V I	•	•	•	•	
λ =	•	•	•	•	
	y(t-2)	y(t-1-n)	u(t-1-k)	u(t-1-k-m)	
	y(t-1)	y(t-n)	u(tlk)	u(t-k-m)	(A2.6)

Now, defining estimated parameters vector $\hat{\underline{\theta}}$ as:

 $\hat{\underline{a}} = (\hat{-a_1}, \hat{-a_2}, \dots, \hat{-a_n}, \hat{b}_0, \hat{b}_1, \dots, \hat{b}_m)$ $= (\hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_N)$

where,



θ₁, θ₂, are the estimated values of University of Moratuwa, Sri Lanka. Feleanous respective Dissertations www.lib.mrt.ac.lk

Y can be expressed as:

$$\underline{Y} = X \frac{\theta}{\theta} + \underline{e}$$
 (A2.7)

where $\underline{e} = (e_1, e_2, \dots, e_n)^T$ is the estimation error vector and e_1, e_2, \dots, e_N are the estimation errors of $y(t-N+1), y(t-N+2), \dots, y(t)$ respectively.

The best estimate for $\underline{\theta}$ is obtained when the scalar product S given by S = $\underline{e}^{T} \underline{e}$ is minimum.

Since $\underline{e} = \underline{Y} - \underline{X} \stackrel{\circ}{\underline{\theta}}$, S can be expressed as:

$$S = \left[\underline{Y} - X\hat{\underline{\theta}}\right]^{T} \left[\underline{Y} - X\hat{\underline{\theta}}\right]$$
(A2.8)

The minimum of S is obtained when $\frac{ds}{d\hat{\theta}}$ is zero.

i,c.

$$\frac{ds}{d\hat{\theta}} = [\underline{Y} - X \hat{\underline{\theta}}]^T X - X^T [\underline{Y} - X \hat{\underline{\theta}}]$$

$$= 2X^T [\underline{Y} - X \hat{\underline{\theta}}]$$

$$= 0$$
Hence, the best $\hat{\underline{\theta}}$ is given by
$$\underline{Y} - X \hat{\underline{\theta}} = 0$$

$$X \hat{\underline{\theta}} = \underline{Y}$$

or

i.e.

 $\hat{\underline{\theta}} = [X^T X]^{-1} X^T \underline{Y}$

 $x^{T} x \hat{\underline{\theta}} = x^{T} \underline{Y}$

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$$\frac{\theta}{\theta} = [X^T X]^{-1} [X^T X] \underline{\theta} = \underline{\theta}$$

Therefore, $\hat{\theta} = \theta$ and correct estimates are obtained.

Also since $\underline{Y} = X \underline{\theta}$

and $\underline{Y} = X\hat{\theta} + \underline{e}$

e is given by e =
$$X \theta - X \hat{\theta}$$

Hence, for $\theta = \hat{\theta}$, e = 0.

A2.2 Recursive least squares algorithm

Consider the equations available at present time t when data samples have been collected over N+1 sample periods:

(A2.9)

$$y(t-N) = -a_1 y(t-N-1) \dots -a_n y(t-N-n) + b_0 u(t-N-k) + \dots + b_m u(t-N-k-m) + b_m u(t-N-k-m)$$

$$(t-1) = -a_1 y(t-2) \dots -a_n y(t-1-n) + b_0 u(t-1-k) + \dots + b_m u(t-1-k-m)$$

$$y(t) = -a_1 y(t-1) \dots -a_n y(t-n) + b_0 u(t-k) + \dots + b_m u(t-k-m)$$

I. matrix form this can be expressed as:

$$\frac{Y}{-t} = X_t \frac{\theta}{-t} t$$
 (A2.10)

Equation (10) Electronic Theses & Dissertations as:

$$\frac{Y}{t} = \begin{bmatrix} \frac{Y}{t-1} \\ \frac{Y}{y(t)} \\ \frac{Y}{t} \end{bmatrix} = \begin{bmatrix} x_{t-1} \\ \frac{-T}{T} \\ \frac{X}{t} \end{bmatrix} = \begin{bmatrix} 0 \\ t \\ \frac{Y}{t-1} \end{bmatrix}$$

$$\frac{Y}{t-1} = \begin{bmatrix} y(t-N), \dots, y(t-2), y(t-1) \end{bmatrix}^{T}$$

$$\frac{x_{t}}{t} = \begin{bmatrix} y(t-1), \dots, y(t-n), u(t-k), \dots, u(t-k-m) \end{bmatrix}$$

where

According to equation (A2.9), the best estimate $\hat{\theta}_t$ is given by:

$$\begin{array}{rcl} & \widehat{\underline{\theta}}_{t} & = & \begin{bmatrix} x_{t}^{T} & x_{t} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t} \end{bmatrix} & \underline{Y}_{t} \\ & & & \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} & \underline{Y}_{t-1} \\ & & & \\ \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} & \underline{X}_{t} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{X}_{t-1} & \underline{X}_{t} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1} & \begin{bmatrix} x_{t-1} \\ \underline{Y}_{t-1} \end{bmatrix}^{-1$$

i.e.
$$\hat{\theta}_{t} = [X_{t-1}^{T} X_{t} + \underline{x}_{t} \underline{x}_{t}^{T}]^{-1} [X_{t-1}^{T} \underline{Y}_{t-1} + \underline{x}_{t} y(t)]$$
(A2.11)

Now, using the matrix identity given by:

$$\begin{bmatrix} A^{T}A + \underline{a} \ \underline{a}^{T} \end{bmatrix}^{-1} = \begin{bmatrix} A^{T}A \end{bmatrix}^{-1} - \underbrace{\begin{bmatrix} A^{T}A \end{bmatrix}^{-1}}_{1 + \underline{a}^{T}} \begin{bmatrix} A^{T}A \end{bmatrix}^{-1} \underline{a}$$

and letting

$$P_{t-1} = \begin{bmatrix} X_{t-1}^{T} & X_{t-1} \end{bmatrix}^{-1}$$
(A2.12)

м

and
$$d_{t-1} = 1 + \frac{x}{t} t^{T} P_{t-1} \frac{x}{t} t$$
 (A2.13)

Equation (A2.11) simplifies to Moratuwa, Sri Lanka.
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$$\vec{v}_t$$
 = $P_{t-1} X_{t-1}^T \underline{Y}_{t-1} + \frac{P_{t-1} \underline{x}_t}{d_{t-1}} [X_{t-1} \underline{Y}_{t-1} + \underline{x}_t y(t)]$
= $P_{t-1} X_{t-1}^T \underline{Y}_{t-1} + \frac{P_{t-1} \underline{x}_t}{d_{t-1}} [d_{t-1} y(t) - \underline{x}_t^T P_{t-1} (X_{t-1}^T \underline{Y}_{t-1} + \underline{x}_t y(t))]$

$$= P_{t-1} X_{t-1}^{T} Y_{t-1} + \frac{P_{t-1} X_{t}}{d_{t-1}} [(1 + X_{t}^{T} P_{t-1} X_{t}) y(t) - X_{t}^{T} P_{t-1} (X_{t-1}^{T} Y_{t-1} + X_{t}) y(t)]$$

$$= P_{t-1} X_{t-1}^{T} Y_{t-1} + \frac{P_{t-1} X_{t}}{d_{t-1}} [y(t) - X_{t}^{T} P_{t-1} X_{t-1}^{T} Y_{t-1}]$$
(A2.14)

According to equation (A2.9) $\hat{\theta}_{t-1}$ is given by:

$$\frac{\hat{\theta}}{\theta_{t-1}} = [X_{t-1}^T \ X_{t-1}]^{-1} \ X_{t-1}^T \ \underline{Y}_{t-1}$$
$$= P_{t-1} \ X_{t-1}^T \ \underline{Y}_{t-1}$$

$$\hat{\underline{\theta}}_{t} = \hat{\underline{\theta}}_{t-1} + \frac{P_{t-1} \underline{x}_{t}}{d_{t-1}} [y(t) - \underline{x}_{t} \hat{\underline{\theta}}_{t-1}]$$
(A2.15)

Since $x = t = \frac{\hat{\theta}}{t-1}$ is the predicted value of y(t) based on the old estimates $\hat{\underline{\theta}}_{t-1}$, the difference $y(t) - \underline{x}_t^T \hat{\underline{\theta}}_{t-1}$ represents the prediction error at time t.

$$\frac{G}{d} \triangleq \frac{P_{t-1} \times t}{d_{t-1}} \text{ is a (n x 1) vector.}$$

Hence equation (A2.15) can be expressed as

$$\hat{\underline{\theta}}_{t} = \hat{\underline{\theta}}_{t-1} + \underline{G} . \text{ (Prediction error)}$$

$$(A2.16)$$

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$$\text{Electronic Theses & Dissertations}_{= \frac{[X + X]}{W + W + 1} \frac{X}{D} - 1} \prod_{t=1}^{T} \frac{[X + X]}{[X + 1]} \prod_{t=1}^{T} \frac{[X + X]}{[X + 1]}$$

i

$$P_{t} = P_{t-1} - \frac{P_{t-1} + x_{t} + x_{t}}{d_{t-1}} P_{t-1}$$
.e.
$$P_{t} = P_{t-1} - G \cdot x_{t} + P_{t-1}$$
(A2.17)

Equations (A2.16) and (A2.17) provide a recursive algorithm for on line parameter estimation. The matrix inversion involved in the definition $P_{t-1} = [X_{t-1}^T X_t]^{-1}$ is avoided as P_t is evaluated in a recursive manner using equation (A2.17).

Appendix 3

Generalised Minimum Variance Control Law

A suitable control u(t), which minimises a general cost function will be derived.

Consider a model given by:

$$y(t) = q^{-k} \frac{B}{A} u(t) + \frac{C}{A} \xi(t)$$
 (A3.1)

and an auxiliary output ϕ defined as:

$$\phi(t+k) \Delta P y(t+k) + Q u(t) - R w(t)$$
(A3.2)

where, P, Q, R are weighting polynomials in

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The control objective is to minimise:

$$I = E \{\phi^2 (t+k)\}$$
(A3.3)

i.e.
$$I = E \{ [P y(t+k) + Q u(t) - R w(t)]^2 \}$$
 (A3.4)

v(t+k) in equation (A3.4) can be expressed in terms of u(t) and ξ (t+k) using equation (A3.1) and the cost function I can be rewritten as:

$$I = E \{ \left[\left(\frac{PB}{A} + Q \right) u(t) - R w(t) + \frac{PC}{A} \xi (t+k) \right]^2 \}$$
(A3.5)

Now, expressing PC/A in the form:

$$\frac{PC}{A} = F + q^{-k} \frac{G}{A}$$

the last term in equation (A3.5), i.e. $\frac{PC}{A} \xi$ (t+k), can be expressed in terms of future values, and present and past values of the random disturbance.

i.e.

ŧ

$$\frac{PC}{A} \xi (t+k) = F \xi (t+k) + \frac{G}{A} \xi (t)$$
 (A3.6)

Since, F is of order k-1 the term F ξ (t+k) involves only future values of the random disturbance.

Now, substituting for
$$\frac{PC}{A} \xi$$
 (t+k) in equation (A3.5):
I = E {[($\frac{PB}{A}$ + Q) u(t) - R w(t) + F ξ (t+k) + $\frac{G}{A} \xi$ (t)]²} (A3.7)

The present and past values of the random disturbance can be calculated from the knowledge of A, B, C, k and the present and past values of u and y using equation (A3.1) as follows:

$$\xi (t) = \frac{A}{C} y(t) - q^{-k} \frac{B}{C} u(t)$$

then, the cost function fin request on IA3s2 modifies to:

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$$I = E \left\{ \left[\left(\frac{PB}{A} + Q \right) u(t) - R w(t) + F\xi(t) + \frac{G}{A} \left(\frac{A}{C} y(t) - q^{-k} \frac{B}{C} u(t) \right) \right]^{2} \right\}$$

$$= E \left\{ \left[\left(\frac{B}{C} \left(\frac{PC}{A} - q^{-k} \frac{G}{A} \right) + Q \right) u(t) + \frac{G}{C} y(t) - R w(t) + F\xi(t+k) \right]^{2} \right\}$$

$$= E \left\{ \left[\frac{1}{C} (H u(t) + G y(t) + E w(t)) + F\xi(t+k) \right]^{2} \right\}$$
(A3.8)

where,

$$H = BF + QC \tag{A3.9}$$

and E = -RC

(A3.10)

Expanding equation (A3.8) results in:

$$I = E \{ \left[\frac{1}{C} (Hu(t) + Gy(t) + Ew(t)) \right]^{2} + \frac{2F}{C} (Hu(t) + Gy(t) + Ew(t)) \xi (t+k) + \left[F\xi (t+k) \right]^{2} \}$$

Since the disturbance ξ (t) is a random uncorrelated zero-mean sequence, the expected value of the middle term will be zero. This is because $F \xi$ (t+k) only involves future values of ξ (t) which are uncorrelated with the present and past values of input, output and reference.

Hence,

I = E { [
$$\frac{1}{C}$$
 (Hu(t) + G y(t) + E w(t))]² + [F \xi (t+k)]² }

and for minimum I;

$$\frac{\partial I}{\partial u(t)} = 0$$

then,

$$\frac{\partial I}{\partial u(t)} = \left\{ \frac{2H}{C} (H u(t) + G y(t) + E w(t)) \right\} = 0$$

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i.e., the control law which minimises the variance of $\phi(t+k)$ is given

by:

H u(t) + G y(t) + E w(t) = 0

Appendix 4

AN ALGORITHM TO SOLVE $AH + q^{-1}BG = CT$

The solution algorithm is considered in the following three steps:

- 1) Transform the equation $AH + q^{-1} BG = CT$ into the form $M\mu = v$ by equating the coefficients of the equal powers of q^{-1} ;
- 2) Perform row operations to diagonalise the matrix M ;
- 3) Compute μ by solving the equations obtained from (2) above.

Step 1 : form M and \underline{v}

Let the orders of the polynomials A, B, C and T to be n_A, n_B, n_C University of Moratuwa, Sri Lanka. and n_T respectively Electronic Theses & Dissertations

Then, from equation (4.63) the orders of the polynomials G and H,

i.e. n_G and n_H , are given by:

 $n_{G} = n_{A} - 1$ $n_{H} = n_{B}$ (h₀ is fixed to unity)

Hence, the total number of unknowns is equal to $n_G + n_H + 1$; i.e. the total number of equations n_{EO} is given by:

$$n_{EQ} = n_G + n_H + 1$$

From equation (4.64) the matrix M is given by:





Step 3:

Computation of G and H parameters

Compute μ_j for $j = n_{EQ}$, $n_{EQ} - 1$, ..., 1

and
$$m_{i+j, i+n_A} = -a_j$$
, for $j = 1, n_A$ (A4.4)

formed as follows:

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 i^{th} element of $\underline{\nu}$ is given by:

$$v_i = -a_i + c_i + t_i + \sum_{k=1}^{n_c} c_k t_{i-k}$$
 (A4.5)

with
$$t_{i-k} = 0$$
 for $i-k > n_T$ and for $i-k < 0$
 $t_{i-k} = 1$ for $i-k = 0$
 $a_i = 0$ for $i > n_A$; $c_i = 0$ for $i > n_c$

Step 2: Diagonalise the matrix M University of Moratuwa, Sri Lanka. (a) University of Moratuwa, Sri Lanka. (b) University of Moratuwa, Sri Lanka. (b) University of Moratuwa, Sri Lanka. (c) University of Moratuwa, Sri Lanka. (c) University of Moratuwa, Sri Lanka.

$$\delta_i = b_i / b_{i-1}, \text{ for } i = 1, n_B$$
(A4.6)

(b) Row operations to set the elements below the diagonal of the first n_A columns to zero. subtract δ_k times ith row from jth row for i = 1, n_A and j=i+1, i+n_B - 1. where, k=i-j

thus,

$${}^{m}_{j,l} = {}^{m}_{j,l} - {}^{m}_{j-k,l} {}^{\delta}_{k}$$
(A4.7)

and

 $\underline{\nu}_{j} = \underline{\nu}_{j} - \underline{\nu}_{j-k} \cdot \delta_{k}$ (A4.8)

Now the equation takes the form:

(a) for $j = n_{EQ}$ $v_j = v_i/m_{jj}$ (A4.12)

(b) for
$$n_A < j < n_{EQ}$$

$$\mu_{j} = \begin{bmatrix} \nu_{j} - \Sigma & m_{j, j+k} & \mu_{j+k} \end{bmatrix} / m_{j, j}$$
(A4.13)

(c) for
$$j < n_A$$

$$\mu_{j} = \begin{bmatrix} \nu_{j} & -\Sigma & m_{j,j+k} & \mu_{j+k} \end{bmatrix} / m_{jj}$$
(A4.14)
$$k = n_{\Delta} - j$$



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